

The Random-Effect Model of Bivariate Meta-Analysis Based on the Difference Mean

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Abstract

In this paper, we establish the random-effect model of bivariate meta-analysis based on the difference mean. Under the assumption of the multilevel normal distribution of the difference mean, we obtain the maximum likelihood estimator $\hat{\mu}_{ML}$ of the difference mean by the method of maximum likelihood estimate. We conclude that the maximum likelihood estimator $\hat{\mu}_{ML}$ of the difference mean is unbiased for the difference mean effect μ . In addition, we get the corresponding covariance matrix and the $(1-\alpha)100\%$ confidence interval for the difference mean effect μ .

Keywords

Meta-Analysis, Maximum Likelihood Estimate, Random-Effect Model

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基于均差估计二变量Meta-分析的随机效应模型

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摘要

本文给出已知协方差矩阵的基于均差估值的二变量Meta-分析随机效应模型, 并通过极大似然估计法给出效应量估计并说明了其无偏性。通过计算估计效应量的协方差矩阵, 给出估计效应分量的 $(1-\alpha)100\%$ 置信区间。

关键词

Meta-分析, 极大似然估计, 随机效应模型

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1. 前言

Meta-分析是一种对同一问题的不同研究进行定量合并的一种方法。该方法是通过考虑每个研究内和研究间的差异性以及研究发表偏倚性得到一个综述结果, 从而解决了单个研究样本量小, 区域、种族差异大等的缺点。因此 Meta-分析的结果对很多领域的科学研究都具有指导意义。针对具体问题的 Meta-分析及其统计方法研究已持续了一个世纪之久。最早是在 1904 年 Karl Pearson 针对伤寒疫苗效果进行了合并研究[1]; 接下来的很多年中对某些零星的问题做了类似的综合[2]。但是在过去的四十年中, 人们逐渐意识到医学实验和临床操作需要基于整体的相关性和可靠证据, Meta-分析的影响得到极大的扩展[3]。仅在 2018 年, 关于健康管理的 Meta-分析发表文章将近 4000 篇。

虽然 Meta-分析在很多领域都有很好的应用, 但对 Meta-分析本身的统计方法研究与其应用研究相比就少了很多。从广泛应用的意义上讲, Meta-分析是以估计量的标准误差为权重的点估计加权平均。随机效应模型的 Meta-分析在各个研究效应量不相等这一假设, 具体是以估计量方差及研究间差异方差和的

逆作为权重得到的加权平均效应量，并可扩展为合并研究间协变量估计[4]。

随着数据科学的快速发展，数据维度和数据量急速增加导致单变量 Meta-分析无法满足现实应用的需求。多变量 Meta-分析有许多优点：第一，可以在一个模型框架下得到所有效应量的估计量；第二，可以通过所有效应量的估计量的协方差矩阵描述多效应量之间的关系。第三，可以获得具有更好统计特性的参数估计量；第四，我们可以获得区别于单变量 Meta-分析的潜在机理原因；第五，多变量分析的方法可以在一定程度上降低纳入文献的偏倚性。

由于多变量 Meta-分析与单变量 Meta-分析比较起来有很多的优点，因此多变量 Meta-分析在各个领域的应用研究是很多的。然而，对于多变量 Meta-分析本身统计方法的研究却很少。1988 年，为了研究教练在 SAT 中的作用，Raudenbush SW 等人通过广义最小二乘法建立了多效应量合并的模型[5]；1993 年，van Houwelingen HC 等人在文献[6]首次给出了基于比值比的二变量 Meta-分析模型，并讨论了该模型的异质性检验和敏感性分析；1996 年，Kalaian HA 等人为了研究教练在 SAT 语言和数学方面的指导效果，在文献[7]建立了基于比值比的各项研究内效应量不同情形下的多元 Meta-分析模型，并通过 REML 方法估计了效应量的协方差矩阵；2002 年，van Houw -elingen HC 等人在广义多元混合线性模型的框架下，通过似然估计方法给出了多元 Meta-回归模型，并将此模型扩展到了非正态分布情形[8]。2008 年，Riley RD 等人在文献[9]中通过极大似然估计法给出了一些特殊相关系数情形下的二变量 Meta-分析的协方差矩阵估计；2008 年，Ritz J 等人在文献[10]中通过极大似然估计和估计方程给出了协方差矩阵已知情形下的多元效应量回归参数，并将此模型应用到肺癌发病率的临床推断中。2010 年，Paul M 等人在文献[11]中通过基于可积嵌套拉普拉斯近似的贝叶斯方法给出了多元 Meta-分析的合并效应量估计，这种方法得到的方差估计偏移量更小且稳定。

本文通过极大似然估计法给出已知协方差矩阵的基于均差估值的二变量 Meta-分析随机效应模型，并通过极大似然估计法给出效应量估计。通过计算估计效应量的协方差矩阵，给出效应分量的 $(1-\alpha)100\%$ 置信区间。

2. 问题描述

在协方差矩阵的基于均差估值的二变量 Meta-分析随机效应模型中，我们假设个体量 X_{ij}^C , X_{ij}^T , X_{i2j}^C 和 X_{i2j}^T 是独立的并且是正态分布，其均值分别为 μ_1^C , μ_1^T , μ_2^C 和 μ_2^T ，每个研究内的方差分别为 σ_{i1}^2 , σ_{i2}^2 。因此 \bar{X}_{i1}^C , \bar{X}_{i1}^T , \bar{X}_{i2}^C 和 \bar{X}_{i2}^T 是独立的且服从正态分布，其均值分别为 μ_1^C , μ_1^T , μ_2^C 和 μ_2^T ，方差分别为 $\frac{\sigma_{i1}^2}{n_{i1}^C}$,

$$\frac{\sigma_{i1}^2}{n_{i1}^T}, \frac{\sigma_{i2}^2}{n_{i2}^C}, \frac{\sigma_{i2}^2}{n_{i2}^T}。$$

由上述假设，均差效应量 D_i 的分布可表示为：

$$D_i = \begin{pmatrix} D_{i1} \\ D_{i2} \end{pmatrix} \sim N \left(\begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}, \begin{pmatrix} \sigma_{i1}^2 w_{i1} + \tau^2 & \rho_0 \sigma_{i1} \sigma_{i2} \sqrt{w_{i1} w_{i2}} + \rho_1 \tau \gamma \\ \rho_0 \sigma_{i1} \sigma_{i2} \sqrt{w_{i1} w_{i2}} + \rho_1 \tau \gamma & \sigma_{i2}^2 w_{i2} + \gamma^2 \end{pmatrix} \right)$$

其中 $\mu_1 = \mu_1^T - \mu_1^C$, $\mu_2 = \mu_2^T - \mu_2^C$ 为均值， $w_{i1} = \frac{1}{n_{i1}^T} + \frac{1}{n_{i1}^C}$, $w_{i2} = \frac{1}{n_{i2}^T} + \frac{1}{n_{i2}^C}$ 及 ρ_0 是非随机的。一般地， μ_1 , μ_2 为未知参数。本文主要内容是通过极大似然估计法来估计 μ_1 , μ_2 。为了方便讨论，记

$S_i = \begin{pmatrix} \sigma_{i1}^2 w_{i1} & \rho_0 \sigma_{i1} \sigma_{i2} \sqrt{w_{i1} w_{i2}} \\ \rho_0 \sigma_{i1} \sigma_{i2} \sqrt{w_{i1} w_{i2}} & \sigma_{i2}^2 w_{i2} \end{pmatrix}$ 为研究内均值效应量的协方差矩阵； $\Sigma = \begin{pmatrix} \tau^2 & \rho_1 \tau \gamma \\ \rho_1 \tau \gamma & \gamma^2 \end{pmatrix}$ 为描述研

究间关系的协方差矩阵。

3. 均值效应量的极大似然估计

根据之前的假设效应量 D_i 服从正态分布，因此关于 μ 的似然函数为：

$$L(\mu, D_i, S_i) = \prod_{i=1}^k \frac{1}{2\pi \det(S_i + \Sigma)^{\frac{1}{2}}} \exp \left\{ -\frac{(D_i - \mu)^T (S_i + \Sigma)^{-1} (D_i - \mu)}{2} \right\}$$

其中 $\mu = \begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}$, $S_i + \Sigma = \begin{pmatrix} \sigma_{i1}^2 \omega_{i1} + \tau^2 & \rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma \\ \rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma & \sigma_{i2}^2 \omega_{i2} + \gamma^2 \end{pmatrix}$ 。其所对应的对数似然函数为：

$$\ln L(\mu, D_i, S_i) = -k \ln(2\pi) - \frac{1}{2} \sum_{i=1}^k \ln[\det(S_i + \Sigma)] - \frac{1}{2} \sum_{i=1}^k (D_i - \mu)^T (S_i + \Sigma)^{-1} (D_i - \mu)$$

定理 3.1: 设效应量 D_i 服从 $N\left(\begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}, \begin{pmatrix} \sigma_{i1}^2 \omega_{i1} + \tau^2 & \rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma \\ \rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma & \sigma_{i2}^2 \omega_{i2} + \gamma^2 \end{pmatrix}\right)$ 的正态分布，

则 μ 的极大似然估计量的分量为：

$$\begin{aligned} \hat{\mu}_1 &= \left\{ \sum_{i=1}^k \sum_{j=1}^k \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{i2}^2 \omega_{i2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \left[(\sigma_{j1}^2 \omega_{j1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) \right. \right. \\ &\quad \left. \left. - (\rho_0 \sigma_{j1} \sigma_{j2} \sqrt{\omega_{j1} \omega_{j2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{j2} \sqrt{\omega_{i1} \omega_{j2}} + \rho_1 \tau \gamma) \right. \right. \\ &\quad \left. \left. \times (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma) \right] \right\}^{-1} \left\{ \sum_{i=1}^k \sum_{j=1}^k \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{i2}^2 \omega_{i2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \right. \\ &\quad \times \left[(\sigma_{j1}^2 \omega_{j1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) - (\rho_0 \sigma_{j1} \sigma_{j2} \sqrt{\omega_{j1} \omega_{j2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \left\{ [(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) \right. \\ &\quad \left. - (\rho_1 \tau \gamma + \rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}})(\rho_0 \sigma_{j1} \sigma_{j2} \sqrt{\omega_{j1} \omega_{j2}} + \rho_1 \tau \gamma)] D_{j1} \right. \\ &\quad \left. \left. - [(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\rho_0 \sigma_{j1} \sigma_{j2} \sqrt{\omega_{j1} \omega_{j2}} + \rho_1 \tau \gamma) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)(\sigma_{j1}^2 \omega_{j1} + \tau^2)] D_{j2} \right\} \right\} \\ \hat{\mu}_2 &= \left\{ \sum_{i=1}^k \sum_{j=1}^k \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{i2}^2 \omega_{i2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \left[(\sigma_{j1}^2 \omega_{j1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) \right. \right. \\ &\quad \left. \left. - (\rho_0 \sigma_{j1} \sigma_{j2} \sqrt{\omega_{j1} \omega_{j2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{j2} \sqrt{\omega_{i1} \omega_{j2}} + \rho_1 \tau \gamma) \right. \right. \\ &\quad \left. \left. \times (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma) \right] \right\}^{-1} \left\{ \sum_{i=1}^k \sum_{j=1}^k \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{i2}^2 \omega_{i2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \right. \\ &\quad \times \left[(\sigma_{j1}^2 \omega_{j1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) - (\rho_0 \sigma_{j1} \sigma_{j2} \sqrt{\omega_{j1} \omega_{j2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \left\{ [(\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma) \right. \\ &\quad \left. \times (\sigma_{j2}^2 \omega_{j2} + \gamma^2) - (\sigma_{i2}^2 \omega_{i2} + \gamma^2)(\rho_0 \sigma_{j1} \sigma_{j2} \sqrt{\omega_{j1} \omega_{j2}} + \rho_1 \tau \gamma)] D_{j1} \right. \\ &\quad \left. \left. - [(\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)(\rho_0 \sigma_{j1} \sigma_{j2} \sqrt{\omega_{j1} \omega_{j2}} + \rho_1 \tau \gamma) - (\sigma_{i2}^2 \omega_{i2} + \gamma^2)(\sigma_{j1}^2 \omega_{j1} + \tau^2)] D_{j2} \right\} \right\} \end{aligned}$$

证明: 因为关于 μ 的似然函数为：

$$L(\boldsymbol{\mu}, \mathbf{D}_i, \mathbf{S}_i) = \prod_{i=1}^k \frac{1}{2\pi \det(\mathbf{S}_i + \Sigma)^{\frac{1}{2}}} \exp \left\{ -\frac{(\mathbf{D}_i - \boldsymbol{\mu})^T (\mathbf{S}_i + \Sigma)^{-1} (\mathbf{D}_i - \boldsymbol{\mu})}{2} \right\}$$

其所对应的对数似然函数为:

$$\begin{aligned} \ln L(\boldsymbol{\mu}, \mathbf{D}_i, \mathbf{S}_i) &= -k \ln(2\pi) - \frac{1}{2} \sum_{i=1}^k \ln[\det(\mathbf{S}_i + \Sigma)] - \frac{1}{2} \sum_{i=1}^k (\mathbf{D}_i - \boldsymbol{\mu})^T (\mathbf{S}_i + \Sigma)^{-1} (\mathbf{D}_i - \boldsymbol{\mu}) \\ &= -\frac{1}{2} \sum_{i=1}^k \ln \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{i2}^2 \omega_{i2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)^2 \right] \\ &\quad - \frac{1}{2} \sum_{i=1}^k \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{i2}^2 \omega_{i2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \\ &\quad \times \left[(\sigma_{i2}^2 \omega_{i2} + \gamma^2)(D_{i1} - \mu_1)^2 - 2(\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma) \right. \\ &\quad \left. \times (D_{i1} - \mu_1)(D_{i2} - \mu_2) + (\sigma_{i1}^2 \omega_{i1} + \tau^2)(D_{i2} - \mu_2)^2 \right] - k \ln(2\pi) \end{aligned}$$

上述对数似然函数对 μ_1 , μ_2 的偏导数为:

$$\begin{aligned} \frac{\partial \ln L(\boldsymbol{\mu}, \mathbf{D}_i, \mathbf{S}_i)}{\partial \mu_1} &= \sum_{j=1}^k \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{i2}^2 \omega_{i2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \\ &\quad \times \left[(\sigma_{i2}^2 \omega_{i2} + \gamma^2)(D_{i1} - \mu_1) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)(D_{i2} - \mu_2) \right] \\ \frac{\partial \ln L(\boldsymbol{\mu}, \mathbf{D}_i, \mathbf{S}_i)}{\partial \mu_2} &= \sum_{j=1}^k \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{i2}^2 \omega_{i2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \\ &\quad \times \left[-(\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)(D_{i1} - \mu_1) + (\sigma_{i1}^2 \omega_{i1} + \tau^2)(D_{i2} - \mu_2) \right] \end{aligned}$$

令 $\frac{\partial \ln L(\boldsymbol{\mu}, \mathbf{D}_i, \mathbf{S}_i)}{\partial \mu_1} = 0, \frac{\partial \ln L(\boldsymbol{\mu}, \mathbf{D}_i, \mathbf{S}_i)}{\partial \mu_2} = 0$ 即得:

$$\begin{aligned} &\sum_{j=1}^k \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{i2}^2 \omega_{i2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \\ &\quad \times \left[(\sigma_{i2}^2 \omega_{i2} + \gamma^2)(D_{i1} - \mu_1) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)(D_{i2} - \mu_2) \right] = 0 \\ &\sum_{j=1}^k \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{i2}^2 \omega_{i2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \\ &\quad \times \left[-(\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)(D_{i1} - \mu_1) + (\sigma_{i1}^2 \omega_{i1} + \tau^2)(D_{i2} - \mu_2) \right] = 0 \end{aligned}$$

整理可得:

$$\begin{aligned} &\sum_{j=1}^k \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{i2}^2 \omega_{i2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)^2 \right]^{-1} (\sigma_{i2}^2 + \gamma^2) \omega_{i2} \mu_1 \\ &\quad - \sum_{j=1}^k \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{i2}^2 \omega_{i2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)^2 \right]^{-1} (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma) \mu_2 \\ &= \sum_{j=1}^k \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{i2}^2 \omega_{i2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)^2 \right]^{-1} (\sigma_{i2}^2 + \gamma^2) \omega_{i2} D_{i1} \\ &\quad - \sum_{j=1}^k \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{i2}^2 \omega_{i2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)^2 \right]^{-1} (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma) D_{i2} \end{aligned}$$

因为 $\mathbf{S}_i + \Sigma = \begin{pmatrix} \sigma_{i1}^2 w_{i1} + \tau^2 & \rho_0 \sigma_{i1} \sigma_{i2} \sqrt{w_{i1} w_{i2}} + \rho_1 \tau \gamma \\ \rho_0 \sigma_{i1} \sigma_{i2} \sqrt{w_{i1} w_{i2}} + \rho_1 \tau \gamma & \sigma_{i2}^2 w_{i2} + \gamma^2 \end{pmatrix}$ 可逆, 且 $(\mathbf{S}_i + \Sigma)^{-1}$ 为:

$$[\det(\mathbf{S}_i + \Sigma)]^{-1} \begin{pmatrix} \sigma_{i2}^2 w_{i2} + \gamma^2 & -(\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{w_{i1} w_{i2}} + \rho_1 \tau \gamma) \\ -(\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{w_{i1} w_{i2}} + \rho_1 \tau \gamma) & \sigma_{i1}^2 w_{i1} + \tau^2 \end{pmatrix}$$

其中 $\det(\mathbf{S}_i + \Sigma) = [(\sigma_{i1}^2 w_{i1} + \tau^2)(\sigma_{i2}^2 w_{i2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{w_{i1} w_{i2}} + \rho_1 \tau \gamma)^2]^{-1}$,

因此上述关于 μ_1, μ_2 的线性方程组的解为:

$$\begin{aligned} \hat{\mu}_1 = & \left\{ \sum_{i=1}^k \sum_{j=1}^k [(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{i2}^2 \omega_{i2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)^2]^{-1} [(\sigma_{j1}^2 \omega_{j1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) \right. \\ & - (\rho_0 \sigma_{j1} \sigma_{j2} \sqrt{\omega_{j1} \omega_{j2}} + \rho_1 \tau \gamma)^2]^{-1} [(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) - (\rho_0 \sigma_{j1} \sigma_{j2} \sqrt{\omega_{j1} \omega_{j2}} + \rho_1 \tau \gamma) \\ & \left. \times (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)] \right\}^{-1} \left\{ \sum_{i=1}^k \sum_{j=1}^k [(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{i2}^2 \omega_{i2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)^2]^{-1} \right. \\ & \times [(\sigma_{j1}^2 \omega_{j1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) - (\rho_0 \sigma_{j1} \sigma_{j2} \sqrt{\omega_{j1} \omega_{j2}} + \rho_1 \tau \gamma)^2]^{-1} \left\{ [(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) \right. \\ & - (\rho_1 \tau \gamma + \rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}})(\rho_0 \sigma_{j1} \sigma_{j2} \sqrt{\omega_{j1} \omega_{j2}} + \rho_1 \tau \gamma)] D_{j1} \\ & \left. - [(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\rho_0 \sigma_{j1} \sigma_{j2} \sqrt{\omega_{j1} \omega_{j2}} + \rho_1 \tau \gamma) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)(\sigma_{j1}^2 \omega_{j1} + \tau^2)] D_{j2} \right\} \\ \hat{\mu}_2 = & \left\{ \sum_{i=1}^k \sum_{j=1}^k [(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{i2}^2 \omega_{i2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)^2]^{-1} [(\sigma_{j1}^2 \omega_{j1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) \right. \\ & - (\rho_0 \sigma_{j1} \sigma_{j2} \sqrt{\omega_{j1} \omega_{j2}} + \rho_1 \tau \gamma)^2]^{-1} [(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) - (\rho_0 \sigma_{j1} \sigma_{j2} \sqrt{\omega_{j1} \omega_{j2}} + \rho_1 \tau \gamma) \\ & \left. \times (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)] \right\}^{-1} \left\{ \sum_{i=1}^k \sum_{j=1}^k [(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{i2}^2 \omega_{i2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)^2]^{-1} \right. \\ & \times [(\sigma_{j1}^2 \omega_{j1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) - (\rho_0 \sigma_{j1} \sigma_{j2} \sqrt{\omega_{j1} \omega_{j2}} + \rho_1 \tau \gamma)^2]^{-1} \left\{ [(\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma) \right. \\ & \times (\sigma_{j2}^2 \omega_{j2} + \gamma^2) - (\sigma_{i2}^2 \omega_{i2} + \gamma^2)(\rho_0 \sigma_{j1} \sigma_{j2} \sqrt{\omega_{j1} \omega_{j2}} + \rho_1 \tau \gamma)] D_{j1} \\ & \left. - [(\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)(\rho_0 \sigma_{j1} \sigma_{j2} \sqrt{\omega_{j1} \omega_{j2}} + \rho_1 \tau \gamma) - (\sigma_{i2}^2 \omega_{i2} + \gamma^2)(\sigma_{j1}^2 \omega_{j1} + \tau^2)] D_{j2} \right\} \end{aligned}$$

注: 由 $(\mathbf{S}_i + \Sigma)^{-1}$ 为: $[\det(\mathbf{S}_i + \Sigma)]^{-1} \begin{pmatrix} \sigma_{i2}^2 w_{i2} + \gamma^2 & -(\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{w_{i1} w_{i2}} + \rho_1 \tau \gamma) \\ -(\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{w_{i1} w_{i2}} + \rho_1 \tau \gamma) & \sigma_{i1}^2 w_{i1} + \tau^2 \end{pmatrix}$,

其中 $\det(\mathbf{S}_i + \Sigma) = [(\sigma_{i1}^2 w_{i1} + \tau^2)(\sigma_{i2}^2 w_{i2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{w_{i1} w_{i2}} + \rho_1 \tau \gamma)^2]^{-1}$,

可以给出效应量 D_i 服从正态分布的合并效应量的矩阵形式为:

$$\hat{\boldsymbol{\mu}}_{ML} = \left[\sum_{i=1}^k (\mathbf{S}_i + \boldsymbol{\Sigma})^{-1} \right]^{-1} \sum_{i=1}^k (\mathbf{S}_i + \boldsymbol{\Sigma})^{-1} \mathbf{D}_i.$$

由上述效应量 D_i 服从正态分布的合并效应量均值估计量分量的具体形式, 我们可以得到合并效应量均值估计量的如下性质:

性质 2.1: 设效应量 D_i 服从 $N\left(\begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}, \begin{pmatrix} \sigma_{i1}^2 \omega_{i1} + \tau^2 & \rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma \\ \rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma & \sigma_{i2}^2 \omega_{i2} + \gamma^2 \end{pmatrix}\right)$ 的正

态分布, 其极大似然估计量为: $\hat{\boldsymbol{\mu}}_{ML} = \left[\sum_{i=1}^k (\mathbf{S}_i + \boldsymbol{\Sigma})^{-1} \right]^{-1} \sum_{i=1}^k (\mathbf{S}_i + \boldsymbol{\Sigma})^{-1} \mathbf{D}_i$, 则 $\hat{\boldsymbol{\mu}}_{ML}$ 对于 $\boldsymbol{\mu}$ 是无偏的。

证明: 要证明估计量的无偏性, 需证等式 $E(\hat{\boldsymbol{\mu}}_{ML}) = \boldsymbol{\mu}$ 成立。根据我们给出的效应量 D_i 服从正态分布的合并效应量均值估计量分量的具体形式, 我们需要证明 $E(\hat{\mu}_1) = \mu_1$, $E(\hat{\mu}_2) = \mu_2$ 。因为 $E(D_{i1}) = \mu_1$, $E(D_{i2}) = \mu_2$, 所以

$$\begin{aligned} E(\hat{\mu}_1) &= E \left\{ \left[\sum_{i=1}^k \sum_{j=1}^k \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{j2} \sqrt{\omega_{i1} \omega_{j2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \left[(\sigma_{j1}^2 \omega_{j1} + \tau^2)(\sigma_{i2}^2 \omega_{i2} + \gamma^2) \right. \right. \right. \\ &\quad \left. \left. - (\rho_0 \sigma_{j1} \sigma_{i2} \sqrt{\omega_{j1} \omega_{i2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{j2} \sqrt{\omega_{i1} \omega_{j2}} + \rho_1 \tau \gamma) \right. \right. \\ &\quad \left. \left. \times (\rho_0 \sigma_{i1} \sigma_{j2} \sqrt{\omega_{i1} \omega_{j2}} + \rho_1 \tau \gamma) \right] \right\}^{-1} \left\{ \sum_{i=1}^k \sum_{j=1}^k \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{j2} \sqrt{\omega_{i1} \omega_{j2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \right. \\ &\quad \times \left[(\sigma_{j1}^2 \omega_{j1} + \tau^2)(\sigma_{i2}^2 \omega_{i2} + \gamma^2) - (\rho_0 \sigma_{j1} \sigma_{i2} \sqrt{\omega_{j1} \omega_{i2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \left\{ [(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) \right. \\ &\quad \left. - (\rho_1 \tau \gamma + \rho_0 \sigma_{i1} \sigma_{j2} \sqrt{\omega_{i1} \omega_{j2}})(\rho_0 \sigma_{j1} \sigma_{i2} \sqrt{\omega_{j1} \omega_{i2}} + \rho_1 \tau \gamma)] D_{j1} \right. \\ &\quad \left. \left. - [(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\rho_0 \sigma_{j1} \sigma_{j2} \sqrt{\omega_{j1} \omega_{j2}} + \rho_1 \tau \gamma) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)(\sigma_{j1}^2 \omega_{j1} + \tau^2)] D_{j2} \right\} \right\} \\ &= \left\{ \sum_{i=1}^k \sum_{j=1}^k \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{j2} \sqrt{\omega_{i1} \omega_{j2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \left[(\sigma_{j1}^2 \omega_{j1} + \tau^2)(\sigma_{i2}^2 \omega_{i2} + \gamma^2) \right. \right. \\ &\quad \left. \left. - (\rho_0 \sigma_{j1} \sigma_{i2} \sqrt{\omega_{j1} \omega_{i2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{j2} \sqrt{\omega_{i1} \omega_{j2}} + \rho_1 \tau \gamma) \right. \right. \\ &\quad \left. \left. \times (\rho_0 \sigma_{i1} \sigma_{j2} \sqrt{\omega_{i1} \omega_{j2}} + \rho_1 \tau \gamma) \right] \right\}^{-1} \left\{ \sum_{i=1}^k \sum_{j=1}^k \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{j2} \sqrt{\omega_{i1} \omega_{j2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \right. \\ &\quad \times \left[(\sigma_{j1}^2 \omega_{j1} + \tau^2)(\sigma_{i2}^2 \omega_{i2} + \gamma^2) - (\rho_0 \sigma_{j1} \sigma_{i2} \sqrt{\omega_{j1} \omega_{i2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \left\{ [(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) \right. \\ &\quad \left. - (\rho_1 \tau \gamma + \rho_0 \sigma_{i1} \sigma_{j2} \sqrt{\omega_{i1} \omega_{j2}})(\rho_0 \sigma_{j1} \sigma_{i2} \sqrt{\omega_{j1} \omega_{i2}} + \rho_1 \tau \gamma)] E(D_{j1}) \right. \\ &\quad \left. \left. - [(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\rho_0 \sigma_{j1} \sigma_{j2} \sqrt{\omega_{j1} \omega_{j2}} + \rho_1 \tau \gamma) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)(\sigma_{j1}^2 \omega_{j1} + \tau^2)] E(D_{j2}) \right\} \right\} \end{aligned}$$

$$\begin{aligned}
&= \left\{ \sum_{i=1}^k \sum_{j=1}^k \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{i2}^2 \omega_{i2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \left[(\sigma_{j1}^2 \omega_{j1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) \right. \right. \\
&\quad \left. \left. - (\rho_0 \sigma_{j1} \sigma_{j2} \sqrt{\omega_{j1} \omega_{j2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{j2} \sqrt{\omega_{i1} \omega_{j2}} + \rho_1 \tau \gamma) \right. \right. \\
&\quad \left. \left. \times (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma) \right] \right\}^{-1} \left\{ \sum_{i=1}^k \sum_{j=1}^k \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{i2}^2 \omega_{i2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \right. \\
&\quad \times \left[(\sigma_{j1}^2 \omega_{j1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) - (\rho_0 \sigma_{j1} \sigma_{j2} \sqrt{\omega_{j1} \omega_{j2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \left\{ (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma) \right. \\
&\quad \times (\sigma_{j2}^2 \omega_{j2} + \gamma^2) - (\sigma_{j2}^2 \omega_{j2} + \gamma^2)(\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma) \left. \right\} \mu_1 \\
&\quad \left. - \left[(\rho_1 \tau \gamma + \rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}})(\rho_0 \sigma_{j1} \sigma_{j2} \sqrt{\omega_{j1} \omega_{j2}} + \rho_1 \tau \gamma) - (\sigma_{i2}^2 \omega_{i2} + \gamma^2)(\sigma_{j1}^2 \omega_{j1} + \tau^2) \right] \mu_2 \right\} \\
&= \left\{ \sum_{i=1}^k \sum_{j=1}^k \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{i2}^2 \omega_{i2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \left[(\sigma_{j1}^2 \omega_{j1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) \right. \right. \\
&\quad \left. \left. - (\rho_0 \sigma_{j1} \sigma_{j2} \sqrt{\omega_{j1} \omega_{j2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{j2} \sqrt{\omega_{i1} \omega_{j2}} + \rho_1 \tau \gamma) \right. \right. \\
&\quad \left. \left. \times (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma) \right] \right\}^{-1} \left\{ \sum_{i=1}^k \sum_{j=1}^k \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{i2}^2 \omega_{i2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \right. \\
&\quad \times \left[(\sigma_{j1}^2 \omega_{j1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) - (\rho_0 \sigma_{j1} \sigma_{j2} \sqrt{\omega_{j1} \omega_{j2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \\
&\quad \left. \times \left[(\sigma_{j2}^2 \omega_{j2} + \gamma^2)(\sigma_{i1}^2 \omega_{i1} + \tau^2) - (\rho_0 \sigma_{j1} \sigma_{j2} \sqrt{\omega_{j1} \omega_{j2}} + \rho_1 \tau \gamma)(\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma) \right] \mu_2 \right\} \\
&= \mu_2
\end{aligned}$$

综上所述, $\hat{\boldsymbol{\mu}}_{ML}$ 对于 $\boldsymbol{\mu}$ 是无偏的。

4. 均值效应量极大似然估计量的协方差矩阵

假设所有的研究都提供所有的均值效应。由多变量统计学的大数定理, 合并效应量均值估计量可以近似为一个多元正态分布, 其对应的协方差矩阵可由下面的定理给出:

定理 4.1: 设效应量设效应量 \mathbf{D}_i 服从 $N\left(\begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}, \begin{pmatrix} \sigma_{i1}^2 \omega_{i1} + \tau^2 & \rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma \\ \rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma & \sigma_{i2}^2 \omega_{i2} + \gamma^2 \end{pmatrix}\right)$ 的正

态分布, 其极大似然估计量为: $\hat{\boldsymbol{\mu}}_{ML} = \left[\sum_{i=1}^k (\mathbf{S}_i + \boldsymbol{\Sigma})^{-1} \right]^{-1} \sum_{i=1}^k (\mathbf{S}_i + \boldsymbol{\Sigma})^{-1} \mathbf{D}_i$, 则 $\hat{\boldsymbol{\mu}}_{ML}$ 所对应的协方差矩阵 $\text{Var}(\hat{\boldsymbol{\mu}}_{ML})$ 为:

$$\begin{aligned}
\text{Var}(\hat{\boldsymbol{\mu}}_{ML})_{(1,1)} &= \left\{ \det \left[\sum_{i=1}^k (\mathbf{S}_i + \boldsymbol{\Sigma})^{-1} \right] \right\}^{-1} \left\{ \sum_{i=1}^k \det [(\mathbf{S}_i + \boldsymbol{\Sigma})]^{-1} (\sigma_{i1}^2 \omega_{i1} + \tau^2) \right\}, \\
\text{Var}(\hat{\boldsymbol{\mu}}_{ML})_{(1,2)} &= \left\{ \det \left[\sum_{i=1}^k (\mathbf{S}_i + \boldsymbol{\Sigma})^{-1} \right] \right\}^{-1} \left\{ \sum_{i=1}^k \det [(\mathbf{S}_i + \boldsymbol{\Sigma})]^{-1} (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma) \right\}, \\
\text{Var}(\hat{\boldsymbol{\mu}}_{ML})_{(2,1)} &= \left\{ \det \left[\sum_{i=1}^k (\mathbf{S}_i + \boldsymbol{\Sigma})^{-1} \right] \right\}^{-1} \left\{ \sum_{i=1}^k \det [(\mathbf{S}_i + \boldsymbol{\Sigma})]^{-1} (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma) \right\},
\end{aligned}$$

$$\text{Var}(\hat{\boldsymbol{\mu}}_{ML})_{(2,2)} = \left\{ \det \left[\sum_{i=1}^k (\mathbf{S}_i + \Sigma)^{-1} \right] \right\}^{-1} \left\{ \sum_{i=1}^k \det [(\mathbf{S}_i + \Sigma)]^{-1} (\sigma_{i2}^2 \omega_{i2} + \gamma^2) \right\}.$$

其中 $\text{Var}(\hat{\boldsymbol{\mu}}_{ML})_{(i,j)}$ 表示矩阵 $\text{Var}(\hat{\boldsymbol{\mu}}_{ML})$ 第 i 行第 j 列的分量,

$$\begin{aligned} & \det \left[\sum_{i=1}^k (\mathbf{S}_i + \Sigma)^{-1} \right] \\ &= \sum_{i=1}^k \sum_{j=1}^k \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{i2}^2 \omega_{i2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \\ & \quad \times \left[(\sigma_{j1}^2 \omega_{j1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) - (\rho_0 \sigma_{j1} \sigma_{j2} \sqrt{\omega_{j1} \omega_{j2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \\ & \quad \times \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{j2}^2 \omega_{j2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{j2} \sqrt{\omega_{i1} \omega_{j2}} + \rho_1 \tau \gamma) \right. \\ & \quad \left. (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma) \right] \\ & \det [(\mathbf{S}_i + \Sigma)^{-1}] = \sum_{j=1}^k \left[(\sigma_{i1}^2 \omega_{i1} + \tau^2)(\sigma_{i2}^2 \omega_{i2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma)^2 \right]. \end{aligned}$$

证明: 根据多元随机变量的知识可知: 如果随机变量 $\mathbf{D} \sim N(\boldsymbol{\mu}, \hat{\Sigma})$, 则 $\text{Var}(\mathbf{MD}) = \mathbf{M} \hat{\Sigma} \mathbf{M}^T$, 其中 \mathbf{T} 表示矩阵的转置. 由 $\mathbf{S}_i = \mathbf{S}_i^T$ 且 $\Sigma^T = \Sigma$, 通过计算可得: $(\mathbf{S}_i + \Sigma)^{-1} = \left[(\mathbf{S}_i + \Sigma)^{-1} \right]^T$,

$$\left[\sum_{i=1}^k (\mathbf{S}_i + \Sigma)^{-1} \right]^{-1} = \left\{ \left[\sum_{i=1}^k (\mathbf{S}_i + \Sigma)^{-1} \right]^{-1} \right\}^T. \text{ 则}$$

$$\begin{aligned} \text{Var}(\hat{\boldsymbol{\mu}}_{ML}) &= \text{Var} \left\{ \left[\sum_{i=1}^k (\mathbf{S}_i + \Sigma)^{-1} \right]^{-1} \sum_{i=1}^k (\mathbf{S}_i + \Sigma)^{-1} D_i \right\} \\ &= \text{Var} \left\{ \sum_{i=1}^k \left\{ \left[\sum_{i=1}^k (\mathbf{S}_i + \Sigma)^{-1} \right]^{-1} (\mathbf{S}_i + \Sigma)^{-1} D_i \right\} \right\} \\ &= \sum_{i=1}^k \text{Var} \left\{ \left[\sum_{i=1}^k (\mathbf{S}_i + \Sigma)^{-1} \right]^{-1} (\mathbf{S}_i + \Sigma)^{-1} D_i \right\} \\ &= \sum_{i=1}^k \left\{ \left[\sum_{i=1}^k (\mathbf{S}_i + \Sigma)^{-1} \right]^{-1} (\mathbf{S}_i + \Sigma)^{-1} (\mathbf{S}_i + \Sigma) (\mathbf{S}_i + \Sigma)^{-1} \right\}^T \left\{ \left[\sum_{i=1}^k (\mathbf{S}_i + \Sigma)^{-1} \right]^{-1} \right\}^T \\ &= \sum_{i=1}^k \left\{ \left[\sum_{i=1}^k (\mathbf{S}_i + \Sigma)^{-1} \right]^{-1} (\mathbf{S}_i + \Sigma)^{-1} (\mathbf{S}_i + \Sigma) (\mathbf{S}_i + \Sigma)^{-1} \left[\sum_{i=1}^k (\mathbf{S}_i + \Sigma)^{-1} \right]^{-1} \right\} \\ &= \sum_{i=1}^k \left\{ \left[\sum_{i=1}^k (\mathbf{S}_i + \Sigma)^{-1} \right]^{-1} (\mathbf{S}_i + \Sigma)^{-1} \left[\sum_{i=1}^k (\mathbf{S}_i + \Sigma)^{-1} \right]^{-1} \right\} \\ &= \left[\sum_{i=1}^k (\mathbf{S}_i + \Sigma)^{-1} \right]^{-1} \left[\sum_{i=1}^k (\mathbf{S}_i + \Sigma)^{-1} \right] \left[\sum_{i=1}^k (\mathbf{S}_i + \Sigma)^{-1} \right]^{-1} \\ &= \left[\sum_{i=1}^k (\mathbf{S}_i + \Sigma)^{-1} \right]^{-1} \end{aligned}$$

$$\text{因为 } (\mathbf{S}_i + \Sigma)^{-1} = \left[\det(\mathbf{S}_i + \Sigma) \right]^{-1} \begin{pmatrix} \sigma_{i2}^2 \omega_{i2} + \gamma^2 & -(\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma) \\ -(\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{\omega_{i1} \omega_{i2}} + \rho_1 \tau \gamma) & \sigma_{i1}^2 \omega_{i1} + \tau^2 \end{pmatrix},$$

其中 $\det(\mathbf{S}_i + \Sigma) = (\sigma_{i1}^2 w_{i1} + \tau^2)(\sigma_{i2}^2 w_{i2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{w_{i1} w_{i2}} + \rho_1 \tau \gamma)^2$, 可得 $\hat{\boldsymbol{\mu}}_{ML}$ 所对应的协方差矩阵 $Var(\hat{\boldsymbol{\mu}}_{ML})$ 为:

$$\begin{aligned} Var(\hat{\boldsymbol{\mu}}_{ML})_{(1,1)} &= \left\{ \det \left[\sum_{i=1}^k (\mathbf{S}_i + \Sigma)^{-1} \right] \right\}^{-1} \left\{ \sum_{i=1}^k \det [(\mathbf{S}_i + \Sigma)]^{-1} (\sigma_{i1}^2 w_{i1} + \tau^2) \right\}, \\ Var(\hat{\boldsymbol{\mu}}_{ML})_{(1,2)} &= \left\{ \det \left[\sum_{i=1}^k (\mathbf{S}_i + \Sigma)^{-1} \right] \right\}^{-1} \left\{ \sum_{i=1}^k \det [(\mathbf{S}_i + \Sigma)]^{-1} (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{w_{i1} w_{i2}} + \rho_1 \tau \gamma) \right\}, \\ Var(\hat{\boldsymbol{\mu}}_{ML})_{(2,1)} &= \left\{ \det \left[\sum_{i=1}^k (\mathbf{S}_i + \Sigma)^{-1} \right] \right\}^{-1} \left\{ \sum_{i=1}^k \det [(\mathbf{S}_i + \Sigma)]^{-1} (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{w_{i1} w_{i2}} + \rho_1 \tau \gamma) \right\}, \\ Var(\hat{\boldsymbol{\mu}}_{ML})_{(2,2)} &= \left\{ \det \left[\sum_{i=1}^k (\mathbf{S}_i + \Sigma)^{-1} \right] \right\}^{-1} \left\{ \sum_{i=1}^k \det [(\mathbf{S}_i + \Sigma)]^{-1} (\sigma_{i2}^2 w_{i2} + \gamma^2) \right\}. \end{aligned}$$

其中

$$\begin{aligned} &\det \left[\sum_{i=1}^k (\mathbf{S}_i + \Sigma)^{-1} \right] \\ &= \sum_{i=1}^k \sum_{j=1}^k \left[(\sigma_{i1}^2 w_{i1} + \tau^2)(\sigma_{j2}^2 w_{j2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{j2} \sqrt{w_{i1} w_{j2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \\ &\quad \times \left[(\sigma_{j1}^2 w_{j1} + \tau^2)(\sigma_{i2}^2 w_{i2} + \gamma^2) - (\rho_0 \sigma_{j1} \sigma_{i2} \sqrt{w_{j1} w_{i2}} + \rho_1 \tau \gamma)^2 \right]^{-1} \\ &\quad \times \left[(\sigma_{i1}^2 w_{i1} + \tau^2)(\sigma_{j2}^2 w_{j2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{j2} \sqrt{w_{i1} w_{j2}} + \rho_1 \tau \gamma) (\rho_0 \sigma_{i1} \sigma_{i2} \sqrt{w_{i1} w_{i2}} + \rho_1 \tau \gamma) \right] \\ &\det [(\mathbf{S}_i + \Sigma)^{-1}] = \sum_{j=1}^k \left[(\sigma_{i1}^2 w_{i1} + \tau^2)(\sigma_{j2}^2 w_{j2} + \gamma^2) - (\rho_0 \sigma_{i1} \sigma_{j2} \sqrt{w_{i1} w_{j2}} + \rho_1 \tau \gamma)^2 \right]. \end{aligned}$$

通过 $\hat{\boldsymbol{\mu}}_{ML}$ 所对应的协方差矩阵 $Var(\hat{\boldsymbol{\mu}}_{ML})$ 可以给出单变量和联合变量的置信区域, 其具体形式为:

性质 3.1: 设效应量 \mathbf{D}_i 服从 $N \left(\begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}, \begin{pmatrix} \sigma_{i1}^2 w_{i1} + \tau^2 & \rho_0 \sigma_{i1} \sigma_{i2} \sqrt{w_{i1} w_{i2}} + \rho_1 \tau \gamma \\ \rho_0 \sigma_{i1} \sigma_{i2} \sqrt{w_{i1} w_{i2}} + \rho_1 \tau \gamma & \sigma_{i2}^2 w_{i2} + \gamma^2 \end{pmatrix} \right)$ 的正

态分布, 其极大似然估计量为: $\hat{\boldsymbol{\mu}}_{ML} = \left[\sum_{i=1}^k (\mathbf{S}_i + \Sigma)^{-1} \right]^{-1} \sum_{i=1}^k (\mathbf{S}_i + \Sigma)^{-1} \mathbf{D}_i$, 则合并均值效应量 $\hat{\boldsymbol{\mu}}_{ML}$ 所对应各个分量的 $(1-\alpha)100\%$ 置信区间为:

$$\left(\hat{\mu}_1 - Z_{\frac{\alpha}{2}} \sqrt{Var(\hat{\boldsymbol{\mu}}_{ML})_{(1,1)}}, \hat{\mu}_1 + Z_{\frac{\alpha}{2}} \sqrt{Var(\hat{\boldsymbol{\mu}}_{ML})_{(1,1)}} \right) \text{ 和 } \left(\hat{\mu}_2 - Z_{\frac{\alpha}{2}} \sqrt{Var(\hat{\boldsymbol{\mu}}_{ML})_{(2,2)}}, \hat{\mu}_2 + Z_{\frac{\alpha}{2}} \sqrt{Var(\hat{\boldsymbol{\mu}}_{ML})_{(2,2)}} \right),$$

其中 $Z_{\frac{\alpha}{2}}$ 表示正态分布的 $\frac{\alpha}{2}$ 分位数, $Var(\hat{\boldsymbol{\mu}}_{ML})_{(i,j)}$ 表示矩阵 $Var(\hat{\boldsymbol{\mu}}_{ML})$ 第 i 行第 j 列的分量。

5. 结束语

本文给出已知协方差矩阵的基于均差估值的二变量 Meta-分析随机效应模型, 并通过极大似然估计法给出效应量估计。通过计算估计效应量的协方差矩阵, 给出效应分量的 $(1-\alpha)100\%$ 置信区间。这些结论对临床医学, 特别是流行病学会有较好的统计学和现实意义。

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