

随机收益下对偶盈余风险模型的推导

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摘 要

本文构建一类含泊松收益到达过程的对偶盈余模型, 假定收益间隔服从指数分布、单次收益为独立同分布的随机变量, 给出盈余过程的数学表达。通过变量代换、微分求导等分析手段推导模型满足的积分微分方程, 刻画盈余过程的演化规律, 为后续破产概率、数值仿真与风险度量分析搭建理论基础。研究结果可为创新企业、资源开采类企业的现金流风险管控提供数理分析工具。

关键词

对偶风险模型, 盈余过程, 泊松过程, 积分微分方程

Derivation of the Dual Surplus Risk Model under Random Gains

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Abstract

This paper constructs a dual surplus model with a Poisson gain arrival process, where the inter-arrival times of gains follow an exponential distribution and individual gains are independent and identically distributed random variables, and provides the mathematical formulation of the surplus process. By means of variable substitution, differential derivation and other analytical methods, the integro-differential equation satisfied by the model is derived to characterize the evolution law of the surplus process, which lays a theoretical foundation for subsequent research on ruin probability, numerical simulation and risk measurement. The research results can provide a mathematical analysis tool for cash flow risk management of innovative enterprises and resource exploitation enterprises.

Keywords

Dual Risk Model, Surplus Process, Poisson Process, Integro-Differential Equation

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1. 引言

风险理论起源于经典 Cramér-Lundberg 索赔盈余模型, 该模型以保费持续流入、随机索赔支出为核心特征, 广泛应用于财产保险、人寿保险风险测算。但现实中存在大量反向现金流经营主体: 高新技术企业长期投入研发成本、矿产企业持续承担开采运维开销、版权企业固定运营支出, 而利润、专利分成、矿产收益为随机正向现金流, 此类场景无法用传统保险盈余模型刻画, 对偶风险模型应运而生。近年来, 对偶风险模型的研究引起了越来越多的关注, 参见[1]-[5]。E. S. Andersen 尝试建立了贴现惩罚函数的微分-积分方程, 推导了折现惩罚函数的解析表达式[6]。

国外学者最先提出对偶风险框架, 初期研究集中于指数收益分布下破产概率显式解, 借助拉普拉斯变换求解积分微分方程[7]; 后续研究拓展至马氏到达、相依收益、带红利策略、扰动布朗运动等拓展模型, 完善了对偶模型的理论体系。

国内研究起步相对较晚, 现有成果多聚焦改进传统对偶模型, 引入利率、税收、再保险、随机支出等因素, 但多数文献收益分布假设较为单一, 对通用分布下模型方程推导过程简化较多, 缺少完整严谨的微分方程推导流程。现有研究多侧重破产概率求解, 对模型基础结构、变量变换、方程推导的系统性梳理不足。

2. 模型介绍

考虑以下对偶模型, 盈余过程定义为

$$U(t) = u - ct + \sum_{i=1}^{N(t)} X_i$$

其中 $u \geq 0$ 表示初始盈余, $c > 0$ 为支出率, 收益序列 $\{X_i\}_{i=1}^{\infty}$ 是独立同分布的随机变量, 其概率密度函数为

f_Y , 累计分布函数为 F_Y , $\sum_{i=1}^{N(t)} X_i$ 表示到达时间 t 前的累积获利。

$N(t) = \max\{k : T_0 + T_1 + \dots + T_k \leq t, T_0 = 0\}$ 是时间 t 前的收益次数, 是 poisson 过程, 参数为 λ , 其中 $\{T_i, i \geq 1\}$ 表示第 $i-1$ 次与第 i 次收益的间隔时间, $\{T_i, i \geq 1\}$ 服从参数为 λ 的指数分布, 且是独立同分布 (i.i.d.) 的随机变量序列; 且定义 $V(u) = E[e^{-\delta t} w(U(T^-), |U(T)|) | U(0) = 0]$, 其中 T 为破产时刻, w 为惩罚函数, δ 为贴现因子。

考虑到对偶风险模型常常用于刻画收益型企业的盈余过程, 然而传统模型中“收益到达间隔与收益额相互独立”的假设难以反映现实中间隔越长, 单次收益越高等典型模型的相依特征, 导致风险度量结果与实际偏差较大, 针对这一问题, 考虑索赔额和时间间隔服从以下相依结构:

$$f_{Y,Y}(t, y) = \lambda e^{-(\lambda+\beta)t} f_1(y) + \lambda (e^{-\lambda t} - e^{-(\lambda+\beta)t}) f_2(y)$$

其中 $f_1(y), f_2(y)$ 为混合分布, 更进一步, 可以得到 X_k 的边际分布:

$$f_{X_k}(x) = M_w(-\beta)f_1(x) + (1 - M_w(-\beta))f_2(x) = \frac{\lambda}{\lambda + \beta}f_1(x) + \frac{\beta}{\lambda + \beta}f_2(x)$$

正安全负载条件可写为:

$$\frac{c}{\lambda} - \frac{\lambda\mu_1 + \beta\mu_2}{\lambda + \beta} > 0$$

其中 μ_1, μ_2 为 $f_1(y), f_2(y)$ 的均值。

3. 对偶风险模型积分微分方程推导

由全概率公式, 可以得到:

$$V(u) = \int_0^{\frac{u}{c}} e^{-\delta t} \left\{ \int_0^{b-u+ct} V(u-ct+y, b) f_{V,Y}(t, y) dy + \int_{b-u+ct}^{\infty} [y+u-ct-b+V(b; b)] f_{V,Y}(t, y) dy \right\} dt$$

代入

$$f_{V,Y}(t, y) = \lambda e^{-(\lambda+\beta)t} f_1(y) + \lambda (e^{-\lambda t} - e^{-(\lambda+\beta)t}) f_2(y)$$

得:

$$\begin{aligned} V(u) &= \int_0^{\frac{u}{c}} e^{-\delta t} \left\{ \int_0^{b-u+ct} V(u-ct+y, b) \left\{ \lambda e^{-(\lambda+\beta)t} f_1(y) + \lambda (e^{-\lambda t} - e^{-(\lambda+\beta)t}) f_2(y) \right\} dy dt \right. \\ &\quad \left. + \int_0^{\frac{u}{c}} e^{-\delta t} \left\{ \int_{b-u+ct}^{\infty} [y+u-ct-b+V(b; b)] \left\{ \lambda e^{-(\lambda+\beta)t} f_1(y) + \lambda (e^{-\lambda t} - e^{-(\lambda+\beta)t}) f_2(y) \right\} dy dt \right. \right. \\ &= \int_0^{\frac{u}{c}} \lambda e^{-(\delta+\lambda)t} \left\{ \int_0^{b-u+ct} V(u-ct+y, b) \left\{ e^{-\beta t} f_1(y) + (1 - e^{-\beta t}) f_2(y) \right\} dy dt \right. \\ &\quad \left. + \int_0^{\frac{u}{c}} e^{-\delta t} \left\{ \int_{b-u+ct}^{\infty} [y+u-ct-b+V(b; b)] \left\{ \lambda e^{-(\lambda+\beta)t} f_1(y) \right\} dy dt \right. \right. \\ &\quad \left. + \int_0^{\frac{u}{c}} e^{-\delta t} \left\{ \int_{b-u+ct}^{\infty} [y+u-ct-b+V(b; b)] \left\{ \lambda (e^{-\lambda t} - e^{-(\lambda+\beta)t}) f_2(y) \right\} dy dt \right. \right. \\ &= \int_0^{\frac{u}{c}} \lambda e^{-(\delta+\lambda+\beta)t} \left\{ \int_0^{b-u+ct} V(u-ct+y, b) \{f_1(y) - f_2(y)\} dy dt \right. \\ &\quad \left. + \int_0^{\frac{u}{c}} \lambda e^{-(\delta+\lambda)t} \int_0^{b-u+ct} V(u-ct+y, b) f_2(y) dy dt \right. \\ &\quad \left. + \int_0^{\frac{u}{c}} e^{-(\delta+\lambda+\beta)t} \left\{ \int_{b-u+ct}^{\infty} [y+u-ct-b+V(b; b)] \{f_1(y) - f_2(y)\} dy dt \right. \right. \\ &\quad \left. + \int_0^{\frac{u}{c}} e^{-(\delta+\lambda)t} \left\{ \int_{b-u+ct}^{\infty} [y+u-ct-b+V(b; b)] f_2(y) dy dt \right. \end{aligned}$$

令 $f_1(y) - f_2(y) = g(y)$, 则有

$$\begin{aligned}
V(u, b) &= \int_0^{\frac{u}{c}} \lambda e^{-(\delta+\lambda+\beta)t} \left\{ \int_0^{b-ut+ct} V(u-ct+y, b) g(y) dy dt \right. \\
&\quad + \int_0^{\frac{u}{c}} \lambda e^{-(\delta+\lambda)t} \int_0^{b-ut+ct} V(u-ct+y, b) f_2(y) dy dt \\
&\quad + \int_0^{\frac{u}{c}} e^{-(\delta+\lambda+\beta)t} \left\{ \int_{b-ut+ct}^{\infty} [y+u-ct-b+V(b; b)] g(y) dy dt \right. \\
&\quad \left. \left. + \int_0^{\frac{u}{c}} e^{-(\delta+\lambda)t} \left\{ \int_{b-ut+ct}^{\infty} [y+u-ct-b+V(b; b)] f_2(y) dy dt \right. \right. \right.
\end{aligned}$$

令 $s = u - ct$, 则 $t = \frac{u-s}{c}$, $dt = -\frac{1}{c} ds$, 则

$$\begin{aligned}
V(u, b) &= \frac{1}{c} \int_0^u \lambda e^{-(\delta+\lambda+\beta)\frac{u-s}{c}} \left\{ \int_0^{b-s} V(s+y, b) g(y) dy ds \right. \\
&\quad + \frac{1}{c} \int_0^u \lambda e^{-(\delta+\lambda)\frac{u-s}{c}} \int_0^{b-s} V(s+y, b) f_2(y) dy ds \\
&\quad + \frac{1}{c} \int_0^u \lambda e^{-(\delta+\lambda+\beta)\frac{u-s}{c}} \left\{ \int_{b-s}^{\infty} [y+s-b+V(b; b)] g(y) dy ds \right. \\
&\quad \left. \left. + \frac{1}{c} \int_0^u e^{-(\delta+\lambda)\frac{u-s}{c}} \left\{ \int_{b-s}^{\infty} [y+s-b+V(b; b)] f_2(y) dy ds \right. \right. \right.
\end{aligned}$$

令

$$\begin{aligned}
W_1(s, b) &= \int_0^{b-s} V(s+y, b) g(y) dy + \int_{b-s}^{\infty} [y+s-b+V(b; b)] g(y) dy \\
W_2(s, b) &= \int_0^{b-s} V(s+y, b) f_2(y) dy + \int_{b-s}^{\infty} [y+s-b+V(b; b)] f_2(y) dy
\end{aligned}$$

则

$$V(u, b) = \frac{1}{c} \int_0^u \lambda e^{-(\delta+\beta+\lambda)\frac{u-s}{c}} W_1(s, b) ds + \frac{1}{c} \int_0^u \lambda e^{-(\delta+\lambda)\frac{u-s}{c}} W_2(s, b) ds$$

求导, 得

$$\begin{aligned}
V'(u, b) &= \frac{\lambda}{c} W_1(u, b) - \frac{\lambda}{c} \frac{\delta+\beta+\lambda}{c} \int_0^u e^{-(\delta+\beta+\lambda)\frac{u-s}{c}} W_1(s, b) ds \\
&\quad + \frac{\lambda}{c} W_2(u, b) - \frac{\lambda}{c} \frac{\delta+\lambda}{c} \int_0^u e^{-(\delta+\lambda)\frac{u-s}{c}} W_2(s, b) ds
\end{aligned}$$

用 $\frac{\delta+\beta+\lambda}{c}$ 乘以 $V(u, b)$, 得

$$\frac{\delta+\beta+\lambda}{c} V(u, b) = \frac{1}{c} \frac{\delta+\beta+\lambda}{c} \int_0^u \lambda e^{-(\delta+\beta+\lambda)\frac{u-s}{c}} W_1(s, b) ds + \frac{1}{c} \frac{\delta+\beta+\lambda}{c} \int_0^u \lambda e^{-(\delta+\lambda)\frac{u-s}{c}} W_2(s, b) ds$$

再加上 $V'(u, b)$, 得

$$\left(\frac{\delta+\beta+\lambda}{c}I+\wp\right)V(u,b)=\frac{\lambda}{c}W_1(u,b)+\frac{\lambda}{c}W_2(u,b)+\frac{\beta\lambda}{c^2}\int_0^u e^{-(\delta+\lambda)\left(\frac{u-s}{c}\right)}W_2(s,b)ds$$

定义

$$\left(\frac{\delta+\beta+\lambda}{c}I+\wp\right)V(u,b)=V_1(u,b)$$

其中

$$V_1(u,b)=\frac{\lambda}{c}W_1(u,b)+\frac{\lambda}{c}W_2(u,b)+\frac{\beta\lambda}{c^2}\int_0^u e^{-(\delta+\lambda)\left(\frac{u-s}{c}\right)}W_2(s,b)ds$$

对 $V_1(u,b)$ 求导得:

$$V_1'(u,b)=\frac{\lambda}{c}W_1'(u,b)+\frac{\lambda}{c}W_2'(u,b)+\frac{\beta\lambda}{c^2}\left(-\frac{\beta+\lambda}{c}\right)\int_0^u e^{-(\delta+\lambda)\left(\frac{u-s}{c}\right)}W_2(s,b)ds$$

用 $\frac{\lambda+\beta}{c}$ 乘以 $V_1(u,b)$, 再结合 $V_1'(u,b)$, 得:

$$\begin{aligned}\left(\frac{\lambda+\beta}{c}I+\wp\right)V_1(u,b) &= \frac{\lambda+\beta}{c}\frac{\lambda}{c}W_1(u,b)+\frac{\lambda+\beta}{c}\frac{\lambda}{c}W_2(u,b) \\ &\quad +\frac{\lambda}{c}W_1'(u,b)+\frac{\lambda}{c}W_2'(u,b)+\frac{\beta\lambda}{c^2}W_2(u,b) \\ \left(\frac{\lambda+\beta}{c}I+\wp\right)\left(\frac{\lambda+\beta+\lambda}{c}I+\wp\right)V(u,b) &= \frac{\lambda+\beta}{c}\frac{\lambda}{c}W_1(u,b)+\frac{\lambda^2+2\beta\lambda}{c}W_2(u,b)+\frac{\lambda}{c}W_1'(u,b)+\frac{\lambda}{c}W_2'(u,b) \\ &= \frac{\lambda}{c}\left(\frac{\lambda+\beta}{c}I+\wp\right)W_1(u,b)+\frac{\lambda}{c}\left(\frac{\lambda+\beta+\lambda}{c}I+\wp\right)W_2(u,b)\end{aligned}$$

4. 结果与讨论

经过全概率分解、变量替换与逐项求导化简后得到的相依收益对偶风险模型积分微分方程, 完整刻画任意初始盈余下贴现惩罚函数的动态演化规律, 其核心数学特征可从算子结构、边界条件、依赖关系、定义域四个维度展开分析:

(1) 算子结构: 微分算子与积分卷积

方程左侧包含一阶线性微分项, 表征固定连续支出对盈余的匀速消耗; 右侧含卷积型积分项, 代表随机泊松收益到达带来的盈余跳跃更新, 本文方程无二阶扩散项, 仅保留一阶微分和单重积分, 说明模型仅存在向上正向收益跳变, 不存在连续随机波动, 贴合高新技术、矿产企业仅阶段性获得大额随机收益、日常成本稳定流出的经营特征。

(2) 突破传统独立假设

传统经典对偶模型中, 收益到达间隔与单次收益额相互独立, 积分核仅为收益边缘密度函数; 本文将间隔与收益的相依联合分布嵌入积分内核, 积分项同时包含到达间隔分布与收益额度的联合概率结构, 使积分算子不再是简单一维卷积, 而是二维相依分布的条件期望形式。该结构赋予方程更强的现实适配性: 能够刻画“收益间隔越长, 单次获利规模越大”的正向相依场景, 弥补经典独立假设与真实企业现金流错配的缺陷。

(3) 齐次线性结构与贴现参数的调节作用

方程整体为线性齐次积分微分系统，贴现因子作为常数系数嵌入微分项，具备线性叠加性质：若已知两组初始盈余对应的惩罚函数解，其线性组合仍为方程解。贴现参数可调节远期随机收益对当前盈余估值的权重，贴现率越大，远期随机收益的积分项权重越低，模型更侧重短期现金流风险；贴现率趋近于 0 时，方程退化为无贴现破产惩罚模型，仅关注长期破产行为。

Andersen 构建的贴现惩罚函数积分微分方程基于收益到达间隔与收益额相互独立的基本假设，积分核仅为收益边缘密度，积分项可拆分为到达强度与收益分布的分离乘积形式。文献[6]方程积分项可分离变量，计算简便，但无法刻画收益与到达间隔的相依关系；本文方程积分核为联合分布，不可分离，形式更复杂，但适配相依现金流场景；Andersen 模型仅适用于收益规模不受等待周期影响的标准化企业，本文拓展后方程适用于研发周期越长、专利收益越高、矿产开采周期越长、营收越高的实体企业；经典独立模型解析解易得，适合理论基础推导；本文方程通用性更强。

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